Dissimilarity of species interaction networks: quantifying the effect of turnover and rewiring

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Despite having established its usefulness in the last ten years, the decomposition of ecological networks in components allowing to measure their β -diversity retains some methodological ambiguities. Notably, how to quantify the relative effect of mechanisms tied to interaction rewiring *vs.* species turnover has been interpreted differently by different authors. In this contribution, I present mathematical arguments and numerical experiments that should (i) establish that the decomposition of networks as it is currently done is indeed fit for purpose, and (ii) provide guidelines to interpret the values of the components tied to turnover and rewiring.

Ecological networks are variable both in time and space (Poisot et al. 2015; Trøjelsgaard & Olesen 2016) -1 this variability motivated the emergence of methodology to compare ecological networks, **including** in a 2 way that meshes with the **core concept for the** comparison of ecological communities, **namely** 3 β -diversity (**Poisot et al. 2012**). The need to understand network variability through partitioning in 4 components equivalent to α , β , and γ diversities is motivated by the prospect to further integrate 5 the analysis of species interactions to the analysis of species compositions. Because species that 6 make up the networks do not react to their environment in the same way, and **because interactions are** 7 only expressed in subsets of the environments in which species co-occurr, the β -diversity of 8 networks may behave in complex ways, and its quantification is likely to be ecologically informative. 9 Poisot et al. (2012) and Canard et al. (2014) have suggested an approach to β -diversity for ecological 10 networks which is based on the comparison of the number of shared and unique links among species 11 within a pair of networks. Their approach differentiates this sharing of links between those 12 established between species occurring in both networks, and those established with at least one 13 unique species. This framework is expressed as the decomposition $\beta_{wn} = \beta_{os} + \beta_{st}$, namely the fact 14 that network dissimilarity (β_{wn}) has a component that can be calculated directly from the dissimilarity of 15 interactions between shared species (β_{os}), and a component that cannot (β_{st}). **Presumably, the value of** 16 these components for a pair of networks can generate insights about the mechanisms involved 17 in dissimilarity. 18

This approach has been widely adopted since its publication, with recent examples using it to understand 19 the effect of fire on pollination systems (Baronio et al. 2021); the impact of rewiring on spatio-temporal 20 network dynamics (Campos-Moreno et al. 2021); the effects of farming on rural and urban landscapes on 21 species interactions (Olsson et al. 2021); the impact of environment gradients on multi-trophic 22 metacommunities (Ohlmann et al. 2018); and as a tool to estimate the sampling completeness of 23 networks (Souza et al. 2021). It has, similarly, received a number of extensions, including the ability to 24 account for interaction strength (Magrach et al. 2017), the ability to handle probabilistic ecological 25 networks (Poisot et al. 2016), and the integration into the Local Contribution to Beta Diversity (Legendre 26 & De Cáceres 2013) approach to understand how environment changes drive network dissimilarity (Poisot 27 et al. 2017). 28

Yet, the precise meaning of β_{st} , namely the importance of species turnover in the overall 30 dissimilarity, has been difficult to capture, and a source of confusion for some practitioners. 31 This is not particularly surprising, as this component of the decomposition responds to unique 32 species introducing their unique interactions both between themselves, and with species that 33 are common to both networks fig. 1. For this reason, it is important to come up with guidelines 34 for the interpretation of this measure, and how to use it to extract ecological insights. 35 Furthermore, much like the definition of β -diversity in all its forms is a contentious topic 36 amongst community ecologists (see e.g. Tuomisto 2010), the β -diversity of networks has been 37 submitted to methodological scrutiny over the years. A synthesis of some criticisms, related to 38 the correct denominator to use to express the proportion of different links, has recently been 39 **published** (Fründ 2021). It argues that the calculation of network dissimilarity terms as **originally** 40 outlined by Poisot et al. (2012) is incorrect, as it can lead to over-estimating the role of interactions 41 between shared species in a network ("rewiring"), and therefore underestimate the importance of species 42 turnover across networks. As mist-understanding either of these quantities can lead to biased 43 inferences about the mechanisms generating network dissimilarity, it is important to assess 44 how the values (notably of β_{os} , and therefore of β_{st}) react to methodological choices. 45 Here, I present a mathematical analysis of the Poisot et al. (2012) method, explain how information 46 about species turnover and link rewiring can be extracted from its decomposition, and conduct 47 numerical experiments to guide the interpretation of the β -diversity values thus obtained (with a specific 48

⁴⁹ **focus on** β_{st} **)**. These numerical experiments establish three core facts. First, the decomposition

- ⁵⁰ adequately captures the relative roles of species turnover and interaction rewiring; **second**, the
- ⁵¹ decomposition responds to differences in network structure (like connectance) as expected;
- 52 finally, the decomposition more accurately captures rewiring than the proposed alternative
- ⁵³ using a different denominator put forth by Fründ (2021).

54 Partitioning network dissimilarity

⁵⁵ The approach to quantifying the difference between pairs of networks established in Poisot *et al.* (2012) is

- ⁵⁶ a simple extension of the overall method by Koleff *et al.* (2003) for species dissimilarity **based** on
- presence-absence data. The objects to compare, X_1 and X_2 , are partitioned into three values,

 $a = |X_1 \cup X_2|, b = |X_2 \setminus X_1|$, and $c = |X_1 \setminus X_2|$, where $|\cdot|$ is the cardinality of set \cdot (the number of elements it contains), and \setminus is the set substraction operation. In the perspective of species composition comparison, X_1 and X_2 are the sets of species in either community, so that if $X_1 = \{x, y, z\}$ and $X_2 = \{v, w, x, y\}$, we have $X_1 \cup X_2 = \{v, w, x, y, z\}, X_1 \cap X_2 = \{x, y\}, X_2 \setminus X_1 = \{v, w\}$, and $X_1 \setminus X_2 = \{z\}$. The core message of Koleff *et al.* (2003) is that the overwheling majority of measures of β -diversity can be re-expressed as functions that operate on the cardinality of these sets – this allows to focus on the number of unique and common elements, as outlined in fig. 1.

65 Re-expressing networks as sets

Applying this framework to networks requires a few additional definitions. Although ecologists tend to 66 think of networks as their adjacency matrix (as is presented in fig. 1), this representation is not optimal 67 to reach a robust understanding of which elements should be counted as part of which set when 68 measuring network dissimilarity. For this reason, we need fall back on the definition of a graph as a pair of 69 sets, wherein $\mathcal{G} = (V, E)$. These two components V and E represent vertices (nodes, species) and edges 70 (interactions), where V is specifically a set containing the vertices of \mathcal{G} , and E is a set of ordered pairs, in 71 which every pair is composed of two elements of V; an element $\{i, j\}$ in E indicates that there is an 72 interaction from species i to species j in the network 9. The adjancency matrix A of this network 73 would therefore have a non-zero entry at A_{ij} . 74

In the context of networks comparison (assuming the networks to compare are \mathcal{M} and \mathcal{N}), we can further decompose the contents of these sets as

$$\mathcal{M} = (V_c \cup V_m, E_c \cup E_{sm} \cup E_{um}),$$

77 and

$$\mathcal{N} = (V_c \cup V_n, E_c \cup E_{sn} \cup E_{un}),$$

78

- ⁷⁹ where V_c is the set of **common species**, V_m and V_n are the species belonging only to network *m* and *n*
- (respectively), E_c are the common edges, and E_{sm} and E_{um} are the interactions unique to k involving,

respectively, only species in V_c , and at least one species from V_m (the same notation applies for the subscript _n).

83 Defining the partitions from networks as sets

⁸⁴ The metaweb (Dunne 2006), which is to say the entire regional species pool and their interaction, can be

defined as $\mathcal{M} \cup \mathcal{N}$ (this operation is commutative), which is to say

$$\mathcal{M} \cup \mathcal{N} = (V_c \cup V_m \cup V_n, E_c \cup E_{sm} \cup E_{um} \cup E_{sn} \cup E_{un}).$$

⁸⁶ This operation gives us an equivalent to γ -diversity for networks, in that the set of vertices contains *all* ⁸⁷ species from the two networks, and the set of edges contains *all* the interactions between these species. If, ⁸⁸ further, we make the usual assumption that only species with at least one interaction are present in the set ⁸⁹ of vertices, then all elements of the set of vertices are present at least once in the set of edges, and the set of ⁹⁰ vertices can be entire reconstructed from the set of edges. Although measures of network β -diversity ⁹¹ operate on interactions (not species), this property is maintained at every decomposition we will describe ⁹² next.

⁹³ We can similarly define the intersection (also commutative) of two networks:

$$\mathcal{M} \cap \mathcal{N} = (V_c, E_c).$$

⁹⁴ The decomposition of β -diversity from Poisot *et al.* (2012) uses these components to measure β_{os}

⁹⁵ ("rewiring"), and β_{wn} (the overall dissimilarity including non-shared species). We can express the

⁹⁶ components *a*, *b*, and *c* of Koleff *et al.* (2003) as the cardinality of the following sets:

Component	а	b	С
β_{os}	E_c	E _{sn}	E _{sm}
β_{wn}	E_c	$E_{sn} \cup E_{un}$	$E_{sm} \cup E_{um}$

⁹⁷ It is fundamental to note that these components can be measured entirely from the

⁹⁸ interactions, and that the number of species in either network are never directly involved.

- ⁹⁹ In the following sections, I present a series of calculations aimed at expressing the values of β_{os} ,
- β_{wn} , and therefore β_{st} as a function of species sharing probability (as a proxy for mechanisms
- ¹⁰¹ generating turnover), and link rewiring probability (as a proxy for mechanisms generating
- ¹⁰² differences in interactions among shared species). These calculations are done using
- ¹⁰³ Symbolics.jl (Gowda et al. 2021), and subsequently transformed in executable code for Julia
- 104 (Bezanson2017JulFre?), used to produce the figures.

105 Quantifying the importance of species turnover

The difference between β_{os} and β_{wn} stems from the species dissimilarity between \mathcal{M} and \mathcal{N} , and it is easier to understand the effect of turnover by picking a dissimilarity measure to work as an exemplar. We will use $\beta = (b + c)/(2a + b + c)$, which in the Koleff *et al.* (2003) framework is (Wilson & Shmida 1984). This measure returns values in [0, 1], with 0 meaning complete similarity, and 1 meaning complete dissimilarity.

Based on a partition between three sets of cardinality a, b, and c,

$$\beta_t = \frac{b+c}{2a+b+c}$$

So as to simplify the notation of the following section, I will introduce a series of new variables. Let $C = |E_c|$ be the number of links that are identical between networks (as a mnemonic, *C* stands for "common"); $R = |E_{sn} \cup E_{sm}|$ be the number of links that are not shared, but only involve shared species (*i.e.* links from $\mathcal{M} \cup \mathcal{N}$ established between species from $\mathcal{M} \cap \mathcal{N}$; as a mnemonic, *R* stands for "rewired"); and $T = |E_{un} \cup E_{um}|$ the number of links that are not shared, and involve at least one unique species (as a mnemonic, *T* stands for "turnover").

There are two important points to note here. First, **as mentionned earlier**, the number or proportion of species that are shared is not involved in the calculation. Second, the connectance of either network is not involved in the calculation. That all links counted in *e.g. U* come from \mathcal{M} , or that they are evenly distributed between \mathcal{M} and \mathcal{N} , has no impact on the result. This is a desirable property of the approach: whatever quantitative value of the components of dissimilarity can be interpreted in the light of the connectance and species turnover *without* any risk of circularity; **indeed**, **I present a numerical**

experiment where connectance varies independently later in this manuscript, reinforcing this point.

- The final component of network dissimilarity in Poisot *et al.* (2012) is β_{st} , *i.e.* the part of β_{wn} that is not
- explained by changes in interactions between shared species (β_{os}), and therefore stems from species

turnover. This fraction is defined as $\beta_{st} = \beta_{wn} - \beta_{os}$.

- ¹²⁹ The expression of β_{st} does not involve a partition into sets that can be plugged into the framework of
- Koleff *et al.* (2003), because the part of \mathcal{M} and \mathcal{N} that are composed of their unique species cannot, by
- definition, share interactions. One could, theoretically, express these as $\mathcal{M} \setminus \mathcal{N} = (V_m, E_{um})$ and
- ¹³² $\mathcal{N} \setminus \mathcal{M} = (V_v, E_{un})$ (note the non-commutativity here), but the dissimilarity between these networks is
- trivially maximal for the measures considered.
- ¹³⁴ Using the β_t measure of dissimilarity, we can re-write (using the notation with A, S, and U)

$$\beta_{os} = \frac{\mathbf{R}}{\mathbf{2C} + \mathbf{R}}$$

135

136 and

$$\beta_{wn} = \frac{\mathbf{R} + \mathbf{T}}{\mathbf{2C} + \mathbf{R} + \mathbf{T}}$$

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Note that β_{os} has the form x/y with x = S and y = 2A + S, and β_{wn} has the form (x + k)/(y + k), with k = U. As long as $k \ge 0$, it is guaranteed that $\beta_{wn} \ge \beta_{os}$, and therefore that $0 \ge \beta_{st} \ge 1$; as *C*, *T*, and *R* are cardinalities of sets, they are necessarily satisfying this condition.

We can get an expression for β_{st} , by bringing β_{os} and β_{wn} to a common denominator and simplifying the numerator:

$$\beta_{st} = \frac{2\mathbf{CT}}{(2\mathbf{C} + \mathbf{R})(2\mathbf{C} + \mathbf{R} + \mathbf{T})}$$

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Note that this value varies in a non-monotonic way with regards to the number of interactions that are

part of the common set of species – this is obvious when developing the denominator into

¹⁴⁶ $4C^2 + R^2 + 4CR + 2CT + RT$. As such, we expect that the value of β_{st} will vary in a hump-shaped way with ¹⁴⁷ the proportion of shared interactions. For this reason, Poisot *et al.* (2012) suggest that $\beta_{st}/\beta wn$ (alt. ¹⁴⁸ $1 - \beta_{os}/\beta_{wn}$) is a better indicator of the *relative* importance of turnover processes on network dissimilarity. ¹⁴⁹ This can be calculated as

$$\frac{\beta_{st}}{\beta_{wn}} = \frac{2\mathbf{CT}}{(2\mathbf{C} + \mathbf{S})(2\mathbf{C} + \mathbf{R} + \mathbf{T})} \times \frac{\mathbf{R} + \mathbf{T}}{2\mathbf{C} + \mathbf{R} + \mathbf{T}}$$

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151 which reduces to

$$\frac{\beta_{st}}{\beta_{wn}} = \frac{2\mathbf{CT}}{(2\mathbf{C} + \mathbf{R})(\mathbf{R} + \mathbf{T})}$$

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The roots of this expression are C = 0 (the turnover of species has no contribution to the difference

between β_{wn} and β_{os} if there are no shared species, and therefore no rewiring), and for T = 0 (the turnover

¹⁵⁵ of species has no contribution if all species are shared).

¹⁵⁶ Quantifying the response of network beta-diversity to souces of variation

157 The relative effect of species turnover and link rewiring

- 158 As the decomposition of beta diversity into sets presented above reveals, the value of the
- ¹⁵⁹ components β_{os} and β_{st} will respond to two family of mechanisms: the probability of sharing a
- ¹⁶⁰ species between the two networks, noted *p*, which will impose bounds on the value of *T*; and the
- ¹⁶¹ probability of an interactions between shared species not being rewired, noted q, which will
- ¹⁶² impose bounds on the value of C. These two probabilities represent, respectively, mechanisms

involved in species turnover and link turnover, as per Poisot *et al.* (2015), and the aim of this

¹⁶⁴ numerical experiment is to describe how these families of processes drive network

165 **dissimilarity**.

173

In order to simplify the calculations, I make the assumptions that the networks have equal species richness (noted S), so that $S_1 = S_2 = S$, and the same connectance (noted ρ), so that $\rho_1 = \rho_2 = \rho$. As a consequence, the two networks have the same number of links $L = \rho \times S_1^2 = \rho \times S_2^2$. The assumption of equal connectance will be relaxed in a subsequent numerical experiment. These simplifications allow to express the size of *C*, *R*, and *T* only as functions of *p* and *q*, as they would all be multiplied by *L*, which can therefore be dropped from the calculation.

The value of C is the proportion of shared species p^2 , as per fig. 1, times the proportion of shared 174 links, q, giving $C = qp^2$. Each network has $r = p^2 - (qp^2)$ rewired links, which leads to 175 $R = 2r = 2p^2(1 - q)$. Finally, we can get the number of unique links in each network t by 176 substracting C + r from the total number of links (which, since we scale everything by L, is 1), 177 yielding $t = 1 - qp^2 - p^2 + qp^2$, which is $t = 1 - p^2$. The total number of unique links due to 178 turnover is $T = 2t = 2(1 - p^2)$. It is important to note that C and R, namely the number of links 179 that are kept or rewired, depends on species sharing (p), as the possible size of the overlap 180 between the two networks does, but the quantity of links that are different due to turnover does 181 not depends on rewiring. 182

183 With the values of C, R, and T, we can write

$$\beta_{\rm os} = \frac{2 {\bf p}^2 (1-{\bf q})}{2 {\bf p}^2 {\bf q} + 2 {\bf p}^2 (1-{\bf q})} = \frac{1-{\bf q}}{{\bf q}+1-{\bf q}} = (1-{\bf q})\,.$$

This is a first noteworthy result: the value of β_{os} , in the ideal scenario of equal links and richness, is the probability of link re-wiring. Because this is true regardless of the value of p(species turnover), this makes β_{os} a strongly ecologically informative component.

187 Similarly, we can write

$$\beta_{wn} = \frac{2p^2(1-q) + 2(1-p^2)}{2p^2q + 2p^2(1-q) + 2(1-p^2)} = \frac{p^2(1-q) + (1-p^2)}{p^2q + p^2(1-q) + (1-p^2)} = 1 - qp^2 \,.$$

The overall dissimilarity responds to q (rewiring) linerarly, and to p quadratically (which is
 expected assuming unipartite networks, in which species are present on both sides).

Expressing β_{os} and β_{wn} as functions of p and q trivializes the search for the expression of β_{st} , which is

$$\beta_{\rm st} = 1 - p^2 q - 1 + q = q \times (1 - p^2)$$

It is worth examining this solution in some detail. β_{st} scales linearly with the probability that a link will *not* be rewired – in other words, in a pair of networks for which rewiring is important (*q* goes to 0), species turnover is going to be a *relatively* less important mechanism to dissimilarity. β_{st} increases when turnover is important (*p* goes to 0), and therefore β_{st} represents a *balance* between species turnover and link rewiring. These three values, as well as β_{st}/β_{wn} , are represented in fig. 2.

198 Sensibility of the decomposition to differences in connectance

The results presented in fig. 2 include the strong assumption that the two networks have equal 199 connectance. Although the range of connectances in nature tends to be very strongly conserved 200 within a system, we can relax this assumption, by letting one network have more interactions 201 than the other. Note that for the sake of notation simplicity, I maintain the constraint that the 202 two networks are equally species rich. Therefore, the sole variation in this numerical 203 experiment is that one network has $L_1 = \rho \times a \times S^2$, and the other network has $L_2 = \rho \times S^2$; in 204 other words, $L_1 = a \times L$ and $L_2 = L$. As one step of the components calculations involves a min 205 operation, I will add the constraint that $L_1 \le L_2$, which is to say $0 < a \le 1$. The value of a is the 206 *ratio* of connectances of the two networks, and the terms S^2 and ρ being shared across all 207 factors, they will be dropped from the calculations. 208

The maximal number of links that can be shared is ap^2 (*i.e.* $min(p^2, ap^2)$), as we cannot share more links than are in the sparsest of the two networks. Of these, q are not rewired, leading to

- $C = aqp^2$. The number of links that are rewired in network 1 is the number of its links between
- shared species minus *C*, *i.e.* $r_1 = ap^2 aqp^2 = ap^2(1 q)$, and similarly $r_2 = p^2 aqp^2 = p^2(1 aq)$,
- leading to $R = r_1 + r_2 = p^2 [a(1-q) + 1]$. Using the same approach, we can get $t_1 = a(1-p^2)$ and
- 214 $t_2 = (1 p^2)$, leading to $T = t_1 + t_2 = (1 p^2)(1 + a)$.
- 215 As in the previous section, we can use these values to write

$$\beta_{\rm os} = 1 - 2 \frac{\rm aq}{1+\rm a} \,,$$

$$\beta_{\rm wn}=1-2\frac{{\rm a}p^2q}{1+{\rm a}}\,,$$

216 and

$$\beta_{\rm st} = 2aq \frac{(1-p^2)(1+a)}{a^2+2a+1}$$

217

[Figure 3 about here.]

The values of these components are visualized in fig. 3. The introduction of the connectance 218 ratio makes these expressions marginally more complex than in the case without differences in 219 connectance, but the noteworthy result remains that in the presence of differences of 220 connectance, the value of β_{os} is still independent from species turnover. In fact, there is an 221 important conclusion to be drawn from this expression. The shared species component is by 222 definition square, meaning that from an actual measurement of β_{os} between two networks for 223 which we know the connectance, noted b_{0s} , we can get the probability of rewiring by 224 reorganizing the terms of $\mathbf{b}_{os} = 1 - 2aq/(1+a)$ as 225

$$\mathbf{q}\approx\frac{(\mathbf{1}-\mathbf{b}_{os})(\mathbf{a}+\mathbf{1})}{2\mathbf{a}}\,,$$

which gives the probability of rewiring as 1 - q; note that this is an *approximation*, as it assumes that the connectances of the entire network and the connectances of the shared components are the same.

229 Does the partition of network dissimilarity needs a new normalization?

One of the arguments put forth in a recent paper by Fründ (2021) is that the decomposition outlined 230 above will overestimate the effect of rewiring; I argue that this is based on a misunderstanding of what β_{st} 231 achieves. It is paramount to clarify that β_{st} is not a direct measure of the importance of turnover: it is a 232 quantification of the relative impact of rewiring to overall dissimilarity, which, all non-turnover 233 mechanisms being accounted for in the decomposition, can be explained by turnover mechanisms. In this 234 section, I present two numerical experiments showing (i) that the β_{os} component is in fact an accurate 235 measure of rewiring, and (ii) that β_{st} captures the consequences of species turnover, and of the 236 interactions brought by unique species. 237

238 Illustrations on arbitrarily small networks are biased

²³⁹ We can re-calculate the illustration of Fründ (2021), wherein a pair of networks with two shared

interactions (C = 2) receive either an interaction in T, in R, or in both:

С	Т	R	β_{os}	β_{wn}	β_{st}	β_{st}/β_{wn}
2	0	0	0	0	0	
2	1	0	1/5	1/5	0	0
2	0	1	0	1/5	1/5	0
2	1	1	1/5	1/3	2/15	2/5

The over-estimation argument hinges on the fact that $\beta_{st} < \beta_{os}$ in the last situation (one interaction as 241 rewiring, one as turnover). Reaching the conclusion of an overestimation from this is based on a 242 mis-interpretation of what β_{st} means. The correct interpretation is that, out of the entire network 243 dissimilarity, only three-fifths are explained by re-wiring. The fact that this fraction is not exactly one-half 244 comes from the fact that the Wilson & Shmida (1984) measure counts shared interactions twice (i.e. it has 245 a 2C term), which over-amplifies the effect of shared interactions as the network is really small. Running 246 the same calculations with C = 10 gives a relative importance of the turnover processes of 47%, and β_{st} 247 goes to 1/2 as C/(T + R) increases. As an additional caveat, the value of β_{st} will depend on the measure of 248 beta-diversity used. Measures that do not count the shared interaction twice are not going to amplify the 249

²⁵⁰ effect of rewiring.

Based on the arguments presented above, I do not think the suggestion of Fründ (2021) to change the 251 denominator of β_{os} makes sense as a default; the strength of the original approach by Poisot *et al.* (2012) is 252 indeed that the effect of turnover is based on a rigorous definition of networks as graphs (as opposed to 253 networks as matrices), in which the induction of vertices from the edgelist being compared gives rise to 254 biologically meaningful denominators. The advantage of this approach is that at no time does the turnover 255 of species itself (or indeed, as shown in many places in this manuscript, the network richness), or the 256 connectance of the network, enter into the calculation of the beta-diversity components. As such, it is 257 possible to use β_{os} and β_{wn} in relationship to these terms, calculated externally (as was recently done by 258 e.g. Higino & Poisot 2021), without creating circularities. 259

²⁶⁰ Therefore the argument of Fründ (2021), whereby the β_{os} component should decrease with ²⁶¹ turnover, and be invariant to connectance, does not hold: the very point of the approach is to ²⁶² provide measures that can be interpreted in the light of connectance and species turnover. ²⁶³ Adopting the perspective developed in the previous section, wherein networks are sets and the measures of β -diversity operates on these sets, highlights the conceptual issue in the Fründ (2021)

alternative normalization: they are using components (namely, interactions) of the networks

²⁶⁶ that are *not* directly part of the two networks being compared.

²⁶⁷ Using an alternative normalization trivializes the results

In this numerical experiment, we reproduce the results in fig. 2, but using the alternative 268 normalization described above. The results are presented in fig. 4. Producing the analytical 269 solutions for the various components, following the expressions for C, T, and R given for fig. 2, 270 yields a similar value for β_{wn} (*i.e.* the two approaches estimate the same value for total 271 dissimiliarity), but different values for β_{st} and β_{os} 272 . Specifically, β_{os} becomes $p^2(1-q)$, which becomes dependent on species turnover. This, from 273 an ecological point of view, makes no sense: the quantification of how much shared species 274 interact in a similar way should not depend on how much species actually overlap. The opposite 275 problem arises for β_{st} , which becomes $1 - p^2$. In short, the relative importance of species 276 turnover is simply species turnover itself, and has no information on interaction dissimilarity. 277 Therefore the core issue of the Fründ (2021) alternative is that, by attempting to fix a non-issue 278 (namely the over-estimate of the importance of re-wiring, which is only true in trivially small 279 networks), it blurs the meaning of β_{os} , and renders β_{st} useless as it is a re-expression of species 280

281 beta-diversity.

282

[Figure 4 about here.]

283 Measuring network beta-diversity: recommendations

²⁸⁴ Based on the numerical experiments and the derivations presented in this paper, we can establish

- ²⁸⁵ a number of recommendations for the measurement and analysis of network dissimilarity.
- First, β_{os} allows to estimate the rate of rewiring, which is an important ecological information to
- have; quantifying it properly can give insights as to how networks differ. Second, β_{st} captures
- ²⁸⁸ both turnover and rewiring mechanisms, but its interpretation is easier to accomplish in the
- context of total network dissimilarity, and therefore β_{st}/β_{wn} should be interpreted more

- ²⁹⁰ thoroughly. Finally, because the alternative denominator from Fründ (2021) removes the
- interesting property of β_{os} (independent estimate of rewiring rate), and trivializes the meaning
- ²⁹² of β_{st} (by turning it into species dissimilarity), there seems to be no valid reason to use it.

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Figure 1: The dissimilarity of two networks (green and orange) of equal richness *S* (this also holds for unequal richness) depends on three families of interactions: those that are unique because of species turnover (in a pale color), those that are unique because of rewiring (in a saturated color), and those that are shared (in black). Assuming that the chance of sharing a species between the two networks is *p*, then there can be at most $p^2 \times S^2$ shared links – for this reason, overall network dissimilarity (β_{wn}) will have a component tied to species turnover, which is β_{st} .

Figure 2: Values of β_{os} , β_{wn} , β_{st} , and β_{st}/β_{wn} as a function of the probability q or sharing a link (*x*-axis), and the probability p of sharing a species (*y*-axis). Larger values indicate *more* dissimilarity, such that for p = q = 1 the dissimilarity as measured by $\beta_{wn} = 0$, and for p = q = 0 the dissimilarity as measured by $\beta_{wn} = 1$. As expected, the relative importance of turnover (β_{st}) is maximal when there is no rewiring, and when turnover increases.







Figure 3: Consequences of changing the ratio of connectances between two equally species-rich networks on the decomposition of network beta-diversity, assuming p = 0.8. Networks with stronger differences in connectance will tend to be more similar, because the differences in number of links becomes extreme enough that the chances of all the links in the sparser network being in the denser network increases.



Figure 4: Reproduction of fig. 2 with the alternative denominators proposed by Fründ (2021).